

Forecasting age distribution of life-table death counts via α -transformation

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- 3 Compositional power α -transformation to model & forecast life-table death counts

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$$S^D = \left\{ (d_1, d_2, \dots, d_D)^\top, \quad d_x > 0, \quad \sum_{x=1}^D d_x = c \right\}$$

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- 4 Oeppen (2008) and Bergeron-Boucher et al. (2017) combine CoDa with PCA to model & forecast life-table death counts

clr transformation

- 1 Transform compositional data from simplex to Euclidean space

$$\mathbf{s} = \{s_x\}_{x=1,\dots,D} = \ln \left\{ \frac{d_x}{(\prod_{x=1}^D d_x)^{1/D}} \right\}$$

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- 2 Extending clr, isometric log-ratio (ilr) is

$$\mathbf{z} = \mathbf{H}\mathbf{s}$$

where $\mathbf{H}: (D-1) \times D$: Helmert matrix

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$$\mathcal{H}_D = \begin{bmatrix} \frac{1}{\sqrt{D}} & \frac{1}{\sqrt{D}} & \frac{1}{\sqrt{D}} & \cdots & \frac{1}{\sqrt{D}} \\ \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} & 0 & \cdots & 0 \\ \frac{1}{\sqrt{6}} & \frac{1}{\sqrt{6}} & -\frac{2}{\sqrt{6}} & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \frac{1}{\sqrt{D(D-1)}} & \frac{1}{\sqrt{D(D-1)}} & \frac{1}{\sqrt{D(D-1)}} & \cdots & -\frac{(D-1)}{\sqrt{D(D-1)}} \end{bmatrix}.$$

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- 3 To remove any redundant dimension

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- 1 No guarantee transformed data are multivariate normal distributed

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Australian data

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Australian data

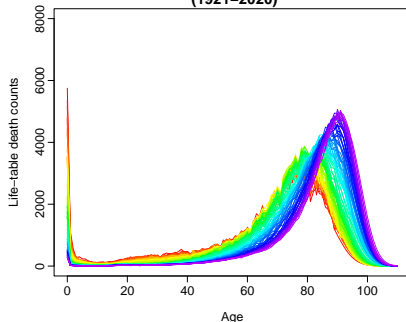
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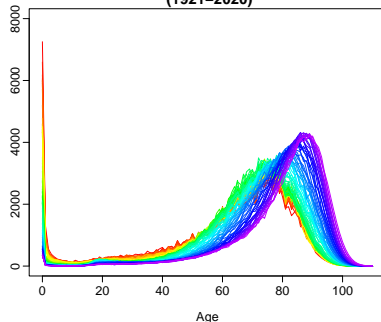
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- 2 Life-table radix is 100,000 at age 0 for each year
- 3 111 ages, ages 0, 1, ..., 109, 110+

Rainbow plot

**Australia: female data
(1921–2020)**



**Australia: male data
(1921–2020)**



- 1 Decreasing trend in infant mortality
- 2 A negatively skewed distribution for life-table death counts
- 3 Shift of the distribution symbolises longevity risk

α transformation

1 α -transformation of a compositional vector $\mathbf{d}_t \in \mathcal{S}^D$

$$\mathbf{z}_t^\alpha = \mathbf{H} \cdot \left(\frac{D\mathbf{d}_t^\alpha - \mathbb{1}_D}{\alpha} \right)$$

$$\mathbf{d}_t^\alpha = \left(\frac{d_{t,1}^\alpha}{\sum_{j=1}^D d_{t,j}^\alpha}, \dots, \frac{d_{t,D}^\alpha}{\sum_{j=1}^D d_{t,j}^\alpha} \right)^\top,$$

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2 α transformation is invertible

$$\mathbf{v}_t^\alpha = D\mathbf{d}_t^\alpha = \alpha \times \mathbf{H}^\top \mathbf{z}_t^\alpha + \mathbf{1}_D$$
$$\mathbf{d}_t = \left(\frac{v_{t,1}^{1/\alpha}}{\sum_{j=1}^D v_{t,j}^{1/\alpha}}, \dots, \frac{v_{t,D}^{1/\alpha}}{\sum_{j=1}^D v_{t,j}^{1/\alpha}} \right)$$

Modeling & forecasting a time series of curves

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- \mathbf{e}_t : model error term

Eigenvalue ratio criterion

- 1 Determine optimal number of components

$$K = \arg \min_{1 \leq k \leq K_{\max}} \left\{ \frac{\hat{\lambda}_{k+1}}{\hat{\lambda}_k} \times \mathbb{1} \left\{ \frac{\hat{\lambda}_k}{\hat{\lambda}_1} \geq \theta \right\} + \mathbb{1} \left\{ \frac{\hat{\lambda}_k}{\hat{\lambda}_1} < \theta \right\} \right\}$$

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$$\hat{z}_{n+h|n}^{\alpha} = E[z_{n+h}^{\alpha} | \hat{\Phi}, z^{\alpha}] = \bar{z}^{\alpha} + \sum_{k=1}^K \hat{\beta}_{n+h|n,k} \hat{\phi}_k$$

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- $\hat{\Phi} = (\hat{\Phi}_1, \dots, \hat{\Phi}_K)^{\top}$

- $\hat{\beta}_{n+h|n,k}$: h -step-ahead forecast of k^{th} principal component scores

Selection of α parameter

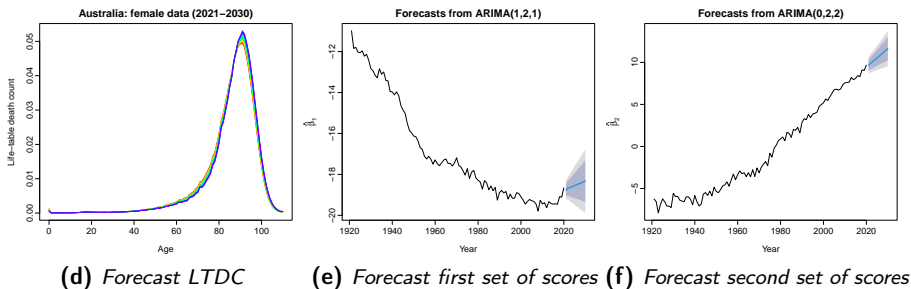
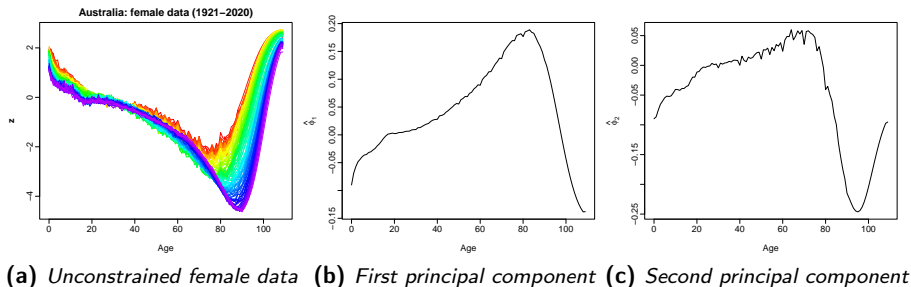
Split data into a training set, a validation set, a testing set

$1 : (n - 20)$

$(n - 19) : (n - 10)$

$(n - 9) : n$

Training	Validation	Testing
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Parameter estimation

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- 3 Multiplying forecast scores with estimated functional principal components to obtain forecast curves

Goodness-of-fit (R^2 criterion)

1 In-sample goodness-of-fit via an R^2 criterion

$$R^2 = 1 - \frac{\sum_{x=1}^{111} \sum_{t=1}^{100} (d_{t,x} - \hat{d}_{t,x})^2}{\sum_{x=1}^{111} \sum_{t=1}^{100} (d_{x,t} - \bar{d}_x)^2}$$

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2 Compared to ilr,

Method	Female	Male
ilr	0.9953	0.9911
α transformation	0.9968	0.9915

Point forecast error measure

- 1 Kullback-Leibler divergence is symmetric & non-negative:

$$\begin{aligned} \text{KLD}(h) &= D_{\text{KL}}(\mathbf{d}_{m+\xi} || \hat{\mathbf{d}}_{m+\xi}) + D_{\text{KL}}(\hat{\mathbf{d}}_{m+\xi} || \mathbf{d}_{m+\xi}) \\ &= \frac{1}{111 \times (11 - h)} \sum_{\xi=h}^{10} \sum_{x=1}^{111} d_{m+\xi,x} \cdot (\ln d_{m+\xi,x} - \ln \hat{d}_{m+\xi,x}) \\ &\quad + \frac{1}{111 \times (11 - h)} \sum_{\xi=h}^{10} \sum_{x=1}^{111} \hat{d}_{m+\xi,x} \cdot (\ln \hat{d}_{m+\xi,x} - \ln d_{m+\xi,x}) \end{aligned}$$

m : end of training or validation set

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- 2 Jensen-Shannon divergence:

$$\text{JSD}(h) = \frac{1}{2} D_{\text{KL}}(\mathbf{d}_{m+\xi} \parallel \boldsymbol{\delta}_{n+\xi}) + \frac{1}{2} D_{\text{KL}}(\hat{\mathbf{d}}_{m+\xi} \parallel \boldsymbol{\delta}_{n+\xi})$$

where $\boldsymbol{\delta}_{m+\xi}$ measures a common quantity between $\mathbf{d}_{m+\xi}$ & $\hat{\mathbf{d}}_{m+\xi}$;

$$\boldsymbol{\delta}_{m+\xi} = \sqrt{\mathbf{d}_{m+\xi} \hat{\mathbf{d}}_{m+\xi}}$$

KLD results

h	Female			Male		
	α -transform	ilr	eda	α -transform	ilr	eda
1	0.0038	0.0078	0.0828	0.0083	0.0107	0.0064
2	0.0040	0.0082	0.0842	0.0106	0.0121	0.0102
3	0.0047	0.0087	0.0853	0.0109	0.0137	0.0175
4	0.0051	0.0087	0.0868	0.0118	0.0151	0.0268
5	0.0054	0.0093	0.0889	0.0152	0.0172	0.0376
6	0.0060	0.0100	0.0928	0.0174	0.0197	0.0511
7	0.0063	0.0101	0.0959	0.0179	0.0230	0.0555
8	0.0082	0.0116	0.0966	0.0246	0.0277	0.0548
9	0.0083	0.0115	0.1043	0.0266	0.0298	0.0786
10	0.0107	0.0132	0.1076	0.0309	0.0345	0.1407
Mean	0.0062	0.0099	0.0925	0.0174	0.0204	0.0479

JSD(geo)

h	Female			Male		
	α -transform	ilr	eda	α -transform	ilr	eda
1	0.0009	0.0019	0.0099	0.0021	0.0027	0.0015
2	0.0010	0.0020	0.0104	0.0024	0.0030	0.0025
3	0.0012	0.0022	0.0108	0.0027	0.0034	0.0044
4	0.0015	0.0022	0.0115	0.0029	0.0038	0.0069
5	0.0013	0.0023	0.0123	0.0033	0.0043	0.0096
6	0.0016	0.0025	0.0133	0.0038	0.0049	0.0131
7	0.0014	0.0025	0.0141	0.0045	0.0057	0.0143
8	0.0024	0.0029	0.0145	0.0061	0.0069	0.0140
9	0.0020	0.0028	0.0165	0.0066	0.0074	0.0203
10	0.0029	0.0033	0.0176	0.0077	0.0086	0.0375
Mean	0.0016	0.0025	0.0131	0.0042	0.0051	0.0124

Interval score

Scoring rule for interval forecasts at time point $d_{n+\xi,x}$:

$$S_{\gamma,\xi}(\hat{d}_{m+\xi,x}^{\text{lb}}, \hat{d}_{m+\xi,x}^{\text{ub}}, d_{m+\xi,x}) = (\hat{d}_{m+\xi,x}^{\text{ub}} - \hat{d}_{m+\xi,x}^{\text{lb}}) + \frac{2}{\gamma}(\hat{d}_{m+\xi,x}^{\text{lb}} - d_{m+\xi,x})\mathbb{1}\{d_{m+\xi,x} < \hat{d}_{m+\xi,x}^{\text{lb}}\} \\ + \frac{2}{\gamma}(d_{m+\xi,x} - \hat{d}_{m+\xi,x}^{\text{ub}})\mathbb{1}\{d_{m+\xi,x} > \hat{d}_{m+\xi,x}^{\text{ub}}\}$$

- γ : level of significance, customarily $\gamma = 0.2$
- For different ages & years in the forecasting period

$$\bar{S}_{\gamma}(h) = \frac{1}{111 \times (11 - h)} \sum_{\xi=h}^{10} \sum_{x=1}^{111} S_{\gamma,\xi}(\hat{d}_{m+\xi,x}^{\text{lb}}, \hat{d}_{m+\xi,x}^{\text{ub}}; d_{m+\xi,x})$$

$$\text{CPD} = |\text{Nominal coverage} - \text{Empirical coverage}|$$

h	Female			Male		
	α -transform	ilr	eda	α -transform	ilr	eda
1	0.0623	0.0759	0.2297	0.0427	0.0759	0.1258
2	0.0655	0.0859	0.2284	0.0819	0.0939	0.1205
3	0.0723	0.0869	0.2189	0.0597	0.0649	0.1315
4	0.0855	0.0958	0.2389	0.0432	0.1060	0.1295
5	0.0934	0.1339	0.2279	0.0357	0.1009	0.1318
6	0.0901	0.1369	0.2432	0.0811	0.1369	0.1445
7	0.0559	0.1640	0.2437	0.1302	0.1392	0.1518
8	0.1129	0.1820	0.2264	0.0829	0.1249	0.2006
9	0.0559	0.1730	0.2279	0.0829	0.1505	0.2910
10	0.1820	0.1820	0.2144	0.1189	0.1369	0.5658
Mean	0.0876	0.1316	0.2300	0.0759	0.1130	0.1993

Mean interval score (\overline{S}_γ)

h	Female			Male		
	α -transform	ilr	eda	α -transform	ilr	eda
1	0.0025	0.0031	0.0051	0.0033	0.0036	0.0024
2	0.0025	0.0031	0.0055	0.0035	0.0036	0.0029
3	0.0028	0.0032	0.0059	0.0037	0.0037	0.0036
4	0.0027	0.0031	0.0063	0.0035	0.0036	0.0041
5	0.0026	0.0032	0.0068	0.0035	0.0036	0.0047
6	0.0029	0.0035	0.0074	0.0037	0.0037	0.0055
7	0.0025	0.0037	0.0079	0.0038	0.0039	0.0064
8	0.0030	0.0036	0.0089	0.0039	0.0040	0.0081
9	0.0028	0.0039	0.0093	0.0043	0.0043	0.0098
10	0.0029	0.0042	0.0103	0.0041	0.0041	0.0137
Mean	0.0027	0.0035	0.0073	0.0037	0.0038	0.0061

Future research

- 1 A robust α transformation

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Future research

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- 2 Extend α transformation to model & forecast life-table death counts for multiple populations
- 3 Consider cohort life-table death counts for specific groups of individuals

Thank you

Paper: <https://arxiv.org/abs/2305.19749>